Assessing disease resistance in chickpeas through the bivariate analysis of Gaussian and binomial traits.

Aidan McGarty¹ Brian Cullis¹, Ahsan Asif² and Kristy Hobson² November 29, 2023

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• Motivation

- Context
- Complexities

• The statistics of the problem

• Choice of model

• Results

• Parametric bootstrapping



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 - Difficult to directly measure/quantify disease presence
 - Field experiments are expensive, labour intensive and time consuming
- These factors lead to a complex statistical analysis



Motivation





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- Hydroponics experiments provide high throughput low cost alternatives to field experiments



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Tank	Rack	HRange	HRow	Variety	TotalLeaf	DisLeaf
1	1	1	1	А	6	2
1	1	2	1	В	5	3
1	1	3	1	С		
1	1	4	1	D	11	6
1	1	5	1	F	13	9
1	1	6	1	А	15	7









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FCol	FRow	Variety	Risk1	Dead1	Risk2	Dead2	Risk3	Dead3	ExpectedLife
1	1	А	16	0	16	12	4	15	1.57
1	2	В	16	0	16	0	16	0	9.39
1	3	С	16	0	16	0	16	2	4.01
1	4	D	10	0	10	7	3	9	1.75
1	5	E	19	0	19	2	17	7	2.98
1	6	А	16	0	16	4	12	7	2.87



Statistics of the Problem



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 - Reflects the plot structure of each experiment
 - Accurately captures the underlying distribution of each measured trait
 - Provides a reliable estimate of the genetic correlation between experiments
- We choose a bivariate generalised linear mixed model (GLMM) using ASRemI-R



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- To overcome this the two data frames are merged by variety in a non-unique way



Statistics of the Problem



WPlot	Variety	FCol	FRow	Tank	Rack	HRange	HRow	DL*	TL*	EL*
1	А	1	1	1	1	1	1	2	6	1.57
2	В	1	2	1	1	2	1	3	5	9.39
3	С	1	3	1	1	3	1			4.01
4	D	1	4	1	1	4	1	6	11	1.75
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*DL = Number of chlorotic leaves (hydroponic), *TL = Number of total leaves (hydroponic) and *EL = Expected lifetime (field)



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• WPlot arbitrarily joins field and hydroponic observations by common varieties, however as evident in WPlot 5 and 6 not all field and hydroponic observations can be matched





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 - y_F = (y_{F1},..., y_{FnF})^T from the field experiment which we assume follows a normal distribution



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- We would also like to include random effects u = (u^T_g, u^T_{pH}, u^T_{pF})^T which are either shared genetic effects u_g or non-shared peripheral effects u_{pH} and u_{pF}, for the hydroponic and field experiments respectively



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- We can then write the distributions $\mathbf{y}_H | \mathbf{u}$ and $\mathbf{y}_F | \mathbf{u}$ conditional on the random effects \mathbf{u} with associated probability density functions (PDF) $f_{Y_H|U}$ and $f_{Y_F|U}$



 $E(y_{H_i}|\mathbf{u}) = \mu_{H_i}$

$$g(\mu_{H_i}) = \eta_{H_i} = \mathbf{x}_{H_i}^{\mathsf{T}} \boldsymbol{ au} + \mathbf{z}_{H_i}^{\mathsf{T}} \mathbf{u}$$

$$V(\mathbf{y}_{H_i}|\mathbf{u}) = \phi_H \mu_{H_i} (1 - \frac{\mu_{H_i}}{n_{TL_i}})$$

- for $i = 1, ..., n_H$ where g() is the logit link function
- x_{Hi} and z_{Hi} are indicator/covariate vectors of length c_X and c_Z relating to the fixed (τ) and random (u) effects for the *i*th observation in the hydroponic experiment
- ϕ_H represents the dispersion parameter and n_{TL_i} are the binomial totals for the *i*th observation in the hydroponic experiment



$$E(y_{F_i}|\mathbf{u}) = \eta_{F_i} = \mathbf{x}_{F_i}^{\mathsf{T}} \boldsymbol{\tau} + \mathbf{z}_{F_i}^{\mathsf{T}} \mathbf{u}$$

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- for $i = 1, \ldots, n_F$
- \mathbf{x}_{F_i} and \mathbf{z}_{F_i} are indicator/covariate vectors of length c_X and c_Z relating to the fixed (τ) and random (\mathbf{u}) effects for the *i*th observation in the field experiment
- σ_F^2 is the field residual variance



Random Effects





$$\begin{bmatrix} \mathbf{u}_g \\ \mathbf{u}_{\rho_H} \\ \mathbf{u}_{\rho_F} \end{bmatrix} \sim N \left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \mathbf{G}_g & \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{G}_{\rho_H} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{G}_{\rho_H} \end{bmatrix} \right)$$



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• Or more generally $\mathbf{u} \sim N(\mathbf{0}, \mathbf{G})$ with associated PDF f_U



$$\begin{bmatrix} \mathbf{u}_g \\ \mathbf{u}_{\mathcal{P}_H} \\ \mathbf{u}_{\mathcal{P}_F} \end{bmatrix} \sim N\left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \mathbf{G}_g & \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{G}_{\mathcal{P}_H} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{G}_{\mathcal{P}_H} \end{bmatrix} \right)$$

- Or more generally $\mathbf{u} \sim N(\mathbf{0}, \mathbf{G})$ with associated PDF f_U
- For convenience we allow κ to contain all parameters in \mathbf{G}_g , \mathbf{G}_{p_H} and \mathbf{G}_{p_F} along with ϕ_H and σ_F^2



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$$egin{aligned} \ell(m{ au},m{\kappa};m{ extbf{y}}) &= \ell(m{ au},m{\kappa};m{ extbf{y}}_H) + \ell(m{ au},m{\kappa};m{ extbf{y}}_F) \ &= \log\left(\int f_{Y_H|U}f_Udm{ extbf{u}}
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$$\ell(\tau, \kappa; \mathbf{y}) = \ell(\tau, \kappa; \mathbf{y}_H) + \ell(\tau, \kappa; \mathbf{y}_F)$$

= $\log\left(\int f_{Y_H|U} f_U d\mathbf{u}\right) + \log\left(\int f_{Y_F|U} f_U d\mathbf{u}\right)$

• The first term of which is not analytically tractable and hence requires an alternative to classical likelihood inference - therefore PQL is used to approximate the likelihood (Collins, 2008)





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• Here $\mathbf{G}_{g_a} = \mathbf{\Sigma}_a \otimes \mathbf{K}$ and $\mathbf{G}_{g_e} = \mathbf{\Sigma}_e \otimes \mathbf{I}_{c_g}$ where \mathbf{K} is a $(c_g \times c_g)$ known genomic relationship matrix (GRM) formed via marker scores



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• Where
$$\Sigma_{a} = \begin{bmatrix} \sigma_{a_{H}} & \sigma_{a_{HF}} \\ \sigma_{a_{HF}} & \sigma_{a_{F}} \end{bmatrix}$$
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- In the case of the non-genetic effects, we specify \mathbf{G}_{p_H} and \mathbf{G}_{p_F} as block diagonal matrices



The Model - Genetic Effects



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• This parameterisation is known as "reduced rank + diag" and is convenient as estimates of the λ_{i_j} , i = a, e, j = H, F cannot go outside the parameter space unlike the " σ " parameters in the unstructured parameterisation



Model - Recap



• Due to the non-Gaussian nature of the response recorded in the hydroponic experiment we use PQL, an approximate likelihood method as the likelihood analytically intractable


- Due to the non-Gaussian nature of the response recorded in the hydroponic experiment we use PQL, an approximate likelihood method as the likelihood analytically intractable
- For the genetic effects, which are our main interest, we specify a variance structure which allows for covariance between experiments



```
asreml(fixed = cbind(DL,EL) ~ trait + trait:VarietyDrop,
random =~ rr(trait):vm(Variety, K) + at(trait, "EL"):vm(Variety, K) +
   rr(trait):ide(Variety) + at(trait, "EL"):ide(Variety) +
   at(trait, "DL"):Tank + at(trait, "DL"):Tank:Rack +
   at(trait, "DL"):Tank:Rack:HRow + at(trait, "DL"):Tank:Rack:HRange +
   at(trait, "EL"):FRow + at(trait, "EL"):FCol,
residual = ~id(WPlot):diag(trait),
 data = df.
family = list(asr_binomial(total = "TL"), asr_gaussian()),
na.action = na.method(x="include",v="include"))
```



Results



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- The researcher would also like a confidence interval on this estimate
- To do so while also assessing potential bias in the PQL estimates, a known issue impacting parameter estimation (Breslow & Lin, 1995) we implement parametric bootstrapping





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- After 50 iterations simulations where $\Delta_{norm} > 0.001$ were excluded





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Parametric Bootstrapping - Sampling Distribution

Parameter	Obs	BootMean
σ^2_{Tank}	0.082	0.077
σ^2_{Rack}	0.022	0.022
σ^2_{FRow}	0.000	0.014
σ^2_{FCol}	0.013	0.025
σ^2_{HRange}	0.038	0.035
$\sigma^2_{\rm HRow}$	0.092	0.088
ψ_{a_F}	0.000	0.023
ψ_{e_F}	2.458	1.617
λ_{a_H}	0.151	0.335
λ_{a_F}	-1.074	-0.032
λ_{e_H}	0.376	0.997
λ_{e_F}	-0.591	-0.874
ϕ_H	1.095	0.965
σ_F^2	3.593	3.571

 σ_j^2 , j = Tank, Rack, FRow, FCol, HRange and HRow are the variance components associated with the peripheral random effects, ψ_{aF} , ψ_{eF} , λ_{aH} , λ_{aF} , λ_{eH} and λ_{eF} are the genetic parameters and ϕ_H and σ_F^2 are the hydroponic dispersion and field residual variance parameter respectively



Conclusion



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- This research question posed a range of complexities to address in order to provide a valid statistical analysis
- The resulting analysis gave a quantification of the level of agreement between PRR resistance for the two experiments
- Parametric bootstrapping provided assessment of parameter estimation bias and indicated the associated 95% bootstrap CI for the total genetic correlation does not contain 0



Rapid and high throughput hydroponics phenotyping method for evaluating chickpea resistance to Phytophthora root rot

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Amritha Amalraj, Julian Taylor, Sean Bithell, Yongle Li, Kevin Moore, Kristy Hobson, and Tim Sutton.

Mapping resistance to phytophthora root rot identifies independent loci from cultivated (cicer arietinum I.) and wild (cicer echinospermum p.h. davis) chickpea.

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